

1 **Title**

2 **Multioomic investigations of Body Mass Index reveal heterogeneous trajectories in response to a**
3 **lifestyle intervention**

4
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17
18 **Abstract**

19 Multiomic profiling is useful in characterizing heterogeneity of both health and disease states. Obesity
20 exerts profound metabolic perturbation in individuals and is a risk factor for multiple chronic diseases.
21 Here, we report a global atlas of cross-sectional and longitudinal changes associated with Body Mass
22 Index (BMI) across 1,100+ blood analytes, as well as their correspondence to host genome and fecal
23 microbiome composition, from a cohort of 1,277 individuals enrolled in a wellness program. Machine
24 learning-based models predicting BMI from blood multiomics captured heterogeneous states of both
25 metabolic and gut microbiome health better than classically measured BMI, suggesting that multiomic
26 data can provide deeper insight into host physiology. Moreover, longitudinal analyses identified
27 variable trajectories of BMI in response to a lifestyle intervention, depending on the analyzed omics
28 platform; metabolomics-based BMI decreased to a greater extent than actual BMI, while proteomics-
29 based BMI exhibited greater resistance. Our analysis further elucidated blood analyte–analyte
30 associations which were significantly modified by obesity and partially reversed in the metabolically
31 obese population through the program. Altogether, our findings provide an atlas of the gradual blood
32 perturbations accompanying obesity and serve as a valuable resource for robustly characterizing
33 metabolic health and identifying actionable targets for obesity.
34

35 Introduction

36 Obesity prevalence has been increasing over the past four decades in adults, adolescents and children
37 around most of the world^{1,2}. Many studies have demonstrated that obesity is a major risk factor for
38 multiple chronic diseases such as type 2 diabetes mellitus (T2DM), metabolic syndrome,
39 cardiovascular disease (CVD), and certain types of cancer³⁻⁶. In obese humans, even 5% loss of body
40 weight can be sufficient for improving metabolic and cardiovascular health⁷, and weight loss with
41 lifestyle interventions can reduce the risk for obesity-related chronic diseases⁸. Nevertheless, obesity
42 and its physiological manifestations appear highly heterogeneous, necessitating additional research to
43 better understand this prevalent health condition.

44 Most commonly, obesity is assessed with an anthropometric Body Mass Index (BMI),
45 defined as the body weight divided by the squared body height [kg m^{-2}]. BMI does not directly
46 measure body composition, but at the population level, BMI correlates well with direct measurements
47 of body fat percentage using computed tomography (CT), magnetic resonance imaging (MRI), or
48 dual-energy X-ray absorptiometry (DXA)⁹. Because BMI is easily calculated and commonly
49 understood among researchers, clinicians, and the general public, it is widely used for the primary
50 diagnosis of obesity and applied as a simple index of efficiency in lifestyle intervention.

51 At the same time, there are considerable limitations to BMI as a surrogate of health/disease
52 metric; e.g., differences in body composition can lead to misclassification of people with a high
53 muscle-to-fat ratio (e.g., athletes) as obese, and can undervalue metabolic improvement with
54 exercise¹⁰. In addition, a meta-analysis demonstrated that the common obesity diagnosis using a BMI
55 cutoff has high specificity but low sensitivity in identifying individuals with excess body fat¹¹, which
56 stems from the difference of appropriate cutoff threshold between ethnic populations as well as the
57 existence of metabolically obese, normal-weight (MONW) individuals^{12,13}. Likewise, there are
58 subgroups in the BMI-classified obese group: i.e., metabolically healthy obese (MHO) and
59 metabolically unhealthy obese (MUO). Although most MHO individuals are not necessarily healthy,
60 and are simply healthier than MUO individuals¹⁴, the transition from MHO to MUO may be an
61 important feature for development of obesity-related chronic diseases¹⁵. Moreover, this transition is
62 potentially preventable through lifestyle interventions¹⁶. Altogether, BMI is undoubtedly useful but
63 too crude to provide insight into heterogeneous states of obesity.

64 Omics studies have suggested the usefulness of blood omics data in health-related metrics;
65 e.g., blood proteomics captures many health conditions¹⁷, and blood metabolomics reflects habitual
66 diets and gut microbiome profiles^{18,19}. Intriguingly, a recent study showed that a machine learning-
67 based BMI model comprising 49 BMI-associated blood metabolites captures obesity-related health
68 outcomes (e.g., percent visceral fat, blood pressure) better than BMI and genetic risk²⁰. Hence,
69 although a targeted metric (e.g., body composition) or a specific biomarker provides useful
70 information, multiomic blood profiling has greater potential to bridge the multifaceted gaps between
71 BMI and the complex physiological states across obese individuals.

72 Previously, our research group has generated personal, dense, dynamic data (PD3) clouds on
73 108 healthy participants as part of a scientific wellness pilot study²¹. The PD3 clouds include human
74 genomes and longitudinal measurements of metabolomics, proteomics, clinical lab tests, gut
75 microbiomes, wearable devices, and health/lifestyle questionnaires accompanied by a coaching
76 intervention. The cohort has since been expanded to over 5,000 individuals, contributing to multiple
77 novel scientific findings^{19,22-26}. In this study, we leveraged the wealth of the collected dataset to
78 investigate the physiological changes accompanying obesity.

79 Herein, using BMI as a starting point, we report our investigations into the multiomic
80 perturbations associated with obesity. Blood analytes across all studied omics platforms have a strong
81 capacity to explain a large portion of the variation in BMI. Through machine learning approaches, we
82 also show a heterogeneity in metabolic states accompanying obesity, which is not captured by
83 measured BMI. Furthermore, longitudinal analyses demonstrate variable changes in metabolic health
84 through lifestyle coaching; i.e., the plasma metabolomics exhibits stronger response than measured

85 BMI, while the plasma proteomics exhibits weaker response. Our findings highlight the power of
86 blood multiomics in investigating the underlying physiology of obesity and weight loss from a clinical
87 standpoint.
88

89 Results

90 Plasma multiomics captures 48–78% of the variance in BMI

91 To investigate the molecular effects of obesity on metabolic profiles, we defined a study cohort of
92 1,277 adults who participated in a wellness program (Arivale)^{19,21–26} and whose datasets included
93 coupled measurements of metabolomics, proteomics, and clinical labs from the same blood draw (see
94 Methods). This study design allowed us to directly investigate the similarities and differences between
95 omics platforms in regards to how they reflected the metabolic state of each individual. The defined
96 cohort was characteristically female (64.3%), middle aged (mean \pm s.d.: 46.6 ± 10.8 years), and white
97 (69.7%) (Supplementary Fig. 1), consistent with our previous studies^{19,22–26}. Based on the World
98 Health Organization (WHO) international standards of BMI cutoff (underweight: $<18.5 \text{ kg m}^{-2}$,
99 normal: $18.5\text{--}25 \text{ kg m}^{-2}$, overweight: $25\text{--}30 \text{ kg m}^{-2}$, obese: $\geq 30 \text{ kg m}^{-2}$)²⁷, the baseline prevalence was
100 similar between normal, overweight, and obese classes, and only 0.8% of participants were
101 underweight class (underweight: 10 participants (0.8%), normal: 426 participants (33.4%),
102 overweight: 391 participants (30.6%), obese: 450 participants (35.2%)).

103 Leveraging the baseline measurements of plasma molecular analytes (metabolomics: 766
104 metabolites, proteomics: 274 proteins, clinical labs: 71 clinical lab tests; Supplementary Data 1), we
105 generated omics-based machine-learning models predicting BMI. To address multicollinearity in our
106 dataset (Supplementary Fig. 2a), we applied a tenfold cross-validation (CV) scheme of the least
107 absolute shrinkage and selection operator (LASSO) algorithm for the baseline metabolomics,
108 proteomics, clinical labs, and combined-omics (i.e., using all metabolomics, proteomics, and clinical
109 labs) datasets. This approach generated metabolomics-based, proteomics-based, clinical labs
110 (chemistries)-based, and combined omics-based BMI predictions (MetBMI, ProtBMI, ChemBMI and
111 CombiBMI, correspondingly). The resulting models retained 62 metabolites, 30 proteins, 20 clinical
112 lab tests, and 132 analytes across all ten MetBMI, ProtBMI, ChemBMI, and CombiBMI models,
113 respectively. These selected predictor variables exhibited low collinearity (Supplementary Fig. 2b, c),
114 validating the variable selection during LASSO modeling²⁸. The generated models demonstrated
115 remarkably high performance for BMI prediction, ranging from mean out-of-sample $R^2 = 0.48$
116 (ChemBMI) to 0.70 (ProtBMI) (Fig. 1a, b). The CombiBMI model further improved the performance
117 of BMI prediction ($R^2 = 0.78$; Fig. 1b), suggesting that, although there is a considerable overlap in the
118 signal detected by each omics platform, different omic measurements still contain unique information
119 regarding BMI. The performance ordering between the models (ChemBMI $<$ ProtBMI \sim MetBMI $<$
120 CombiBMI) were also consistent in sex-stratified models (Supplementary Fig. 3a).

121 BMI has been reported to be associated with many anthropometric and clinical metrics such
122 as waist circumference, blood pressure, sleep quality, and polygenic risk scores (PRSs)^{3,4,14,24,29}. As a
123 first test for the validity of omics-based BMI models, we examined the association of omics-based
124 BMI with each of 51 numeric physiological measures. BMI was significantly associated with 27
125 features (false discovery rate (FDR) < 0.05) including daily physical activity measures from wearable
126 devices, waist-to-height ratio, blood pressure, and BMI PRS (Fig. 1c). With minor differences in
127 significant features between the models (MetBMI: 25 features, ProtBMI: 25 features, ChemBMI: 25
128 features, CombiBMI: 25 features), there was concordance among the associations of all omics-based
129 BMI predictions and these BMI-associated features (Fig. 1c), indicating that the omics-based BMIs
130 basically maintain the characteristics of BMI in terms of anthropometric, genetic, lifestyle, and
131 physiological parameters.
132

133 Omics-based BMI captures the variation in BMI better than any single analyte

134 To confirm the robustness of the variable selection process, we iterated the LASSO modeling while
135 dropping the analyte with the strongest β -coefficient in each iteration step. If a variable is
136 indispensable for a model, the performance should decrease drastically after dropping the variable. In
137 all omics-based BMI models, a steep decay in the mean out-of-sample R^2 across ten models was
138 observed in the first 5–9 iterations (Supplementary Fig. 3b–e), suggesting that the variables that had
139 the largest absolute β -coefficient values in the original LASSO models were the most important in
140 predicting BMI. Interestingly, the overall slope of R^2 decay in the MetBMI model was more gradual
141 than that in the ProtBMI and ChemBMI models (Supplementary Fig. 3b–d), implying that metabolites
142 contain more redundant information to predict BMI. Indeed, the proportion of the variables that were
143 robustly retained across all ten LASSO models to the variables that were retained in at least one of ten
144 LASSO models was lower in the MetBMI model compared to the ProtBMI and ChemBMI models
145 (MetBMI: 62/209 metabolites \approx 29.7%, ProtBMI: 30/74 proteins \approx 40.5%, ChemBMI: 20/41 clinical
146 lab tests \approx 48.9%), confirming the higher level of redundancy in metabolomics. Nevertheless, a large
147 number of metabolites remained in the robust 132 analytes of the CombiBMI model (77 metabolites,
148 51 proteins, 4 clinical lab tests; Fig. 2a), suggesting that each of the data types possesses unique
149 information about BMI. The strongest positive predictors in the CombiBMI model (mean β -coefficient
150 > 0.02) were leptin (LEP), adrenomedullin (ADM), and fatty acid-binding protein 4 (FABP4), and the
151 strongest negative ones (mean β -coefficient < -0.02) were insulin-like growth factor-binding protein 1
152 (IGFBP1) and advanced glycosylation end-product specific receptor (AGER; also described as
153 receptor of AGE, RAGE); that is, proteins were the strongest contributors to the model. Furthermore,
154 although it is possible that metabolites which are highly associated with the retained proteins were
155 eliminated from the CombiBMI model due to collinearity, the absolute β -coefficient values of the
156 robustly retained variables were still lower and the differences among them were milder in the
157 MetBMI model compared to the ProtBMI model (Supplementary Fig. 4).

158 At the same time, the existence of strong and robust predictors in the omics-based BMI model
159 implied that a single analyte may carry sufficient information to predict BMI. To address this
160 possibility, we regressed BMI independently to each analyte that was retained in at least one of ten
161 LASSO models (MetBMI: 209 metabolites, ProtBMI: 74 proteins, ChemBMI: 41 clinical lab tests).
162 Among the analytes that were significantly associated with BMI (180 metabolites, 63 proteins, 30
163 clinical lab tests), only LEP, FABP4, and interleukin 1 receptor antagonist (IL1RN) univariately
164 explained over 30% of the variance in BMI (Fig. 2b–d), with a maximum of 37.9% (LEP). In contrast,
165 the MetBMI, ProtBMI, and ChemBMI models explained 68.9%, 70.6%, and 48.8% of the variance in
166 BMI, respectively. Moreover, even upon eliminating several strong predictor analytes such as LEP
167 and FABP4 from the omic datasets, the generated models still explained larger variance in BMI than
168 any single analyte (Supplementary Fig. 3b–e). These results indicate that the omics-based BMI models
169 explain a larger portion of the variation in BMI than any single analyte.
170

171 **Metabolic heterogeneity within standard BMI classes underlies the high rate of misclassification**

172 Although the omics-based BMIs concordantly represented the characteristics of BMI (Fig. 1c), we still
173 observed that the difference between the measured and predicted BMIs was significantly correlated
174 between the omics-based BMI models (Fig. 3a), implying the cases where the omics-based predictions
175 deviated from the measured BMI were in fact a result of a different underlying metabolic state
176 consistently reflected across the omics-based BMIs, rather than an artifact of the model generation. In
177 addition, upon classifying the participants into the WHO international standard BMI classes based on
178 either the measured or the omics-based BMI values, the misclassification rate was approximately 30%
179 across all omics categories and BMI classes (Fig. 3b), consistent with the previously reported
180 misclassification rates derived from the cardiometabolic health classification^{30,31}.

181 We then re-examined the BMI-associated features (Fig. 1c) while stratifying by measured
182 BMI class and misclassification status; i.e., each participant was classified using both measured BMI
183 and the predicted BMI (MetBMI or ProtBMI) based on the standard BMI cutoffs, and categorized into
184 “Matched” or “Misclassified” when the BMI-based class matched or mismatched the omics-based

185 BMI class, respectively. The misclassified group of normal BMI class individuals exhibited
186 significantly higher values of features that are positively associated with BMI, such as waist-to-height
187 ratio, heart rate, and blood pressure, and significantly lower values of features that are negatively
188 associated with BMI, such as daily physical activity measures, compared to the corresponding
189 matched group of normal BMI class individuals (Fig. 3c), suggesting that a participant misclassified
190 into the normal BMI class possesses an unhealthier molecular profile reflected by metabolomics and
191 proteomics, similar to that of overweight and obese individuals, and corresponding to a MONW
192 individual. Conversely, the misclassified group of obese BMI individuals exhibited significantly lower
193 values of features that are positively associated with BMI and significantly higher values of features
194 that are negatively associated with BMI compared to the corresponding matched group of obese BMI
195 class individuals (Fig. 3c), suggesting that a participant misclassified into the obese BMI class has a
196 healthier metabolomic and proteomic signature, similar to that of overweight and normal individuals,
197 and corresponding to a MHO individual. Importantly, there was no difference in BMI PRS between
198 the matched and misclassified groups (Fig. 3c), implying that the discordance between measured and
199 omics-based BMIs may stem from lifestyle or environmental factors, rather than genetic propensity. In
200 this analysis, although a statistical difference in age between the matched and misclassified groups of
201 the normal BMI class was also observed (Fig. 3c), the age difference does not explain the above
202 differences, as the statistical models were adjusted for age. The findings of concordant patterns
203 between matched and misclassified groups of the normal and obese BMI classes were strengthened by
204 consistent trends in multiple other known obesity-related health markers^{3,14,32-34}, including
205 triglyceride, high-density lipoprotein (HDL) cholesterol, adiponectin, high-sensitivity C-reactive
206 protein (CRP), homeostatic model assessment for insulin resistance (HOMA-IR), glycohemoglobin
207 (HbA1c), and vitamin D (Supplementary Fig. 5a). Taken together, these results suggest that the
208 omics-based BMI models capture the heterogeneous metabolic health states of individuals which are
209 not captured by standard (measured) BMI cutoffs.

210 To further explore the molecule-level difference between the matched and misclassified
211 groups, we applied unsupervised hierarchical clustering using proteomics and metabolomics data. We
212 observed three clusters for the normal BMI class in a proteomic space defined with the strongest 15
213 proteins in the ProtBMI models, and Cluster 3 and Cluster 2 were significantly enriched for the
214 matched and misclassified group individuals, respectively (Fig. 3d). In Cluster 3, the expression levels
215 of these proteins prominently corresponded to their contributing directions in the ProtBMI model. In
216 contrast, this correspondence was blurred in Cluster 2, implying that the misclassification may emerge
217 from dysregulation of strong predictor proteins. In addition, we observed three clusters for the obese
218 BMI class in the same proteomic space, and Clusters 2 and 3 and Cluster 1 were significantly enriched
219 for the matched and misclassified group individuals, respectively (Fig. 3e). Likewise, the strong
220 predictor proteins in Cluster 1 (i.e., the misclassified individuals-enriched cluster in obese BMI class)
221 exhibited weaker agreement between their expression levels and their contributing directions in the
222 ProtBMI model. Furthermore, similar patterns were observed in a metabolomic space, clustering the
223 strongest 15 metabolites in the MetBMI models (Supplementary Fig. 5b, c). These results imply that
224 the conventional BMI classification failed to capture the differences in molecular regulatory states of
225 blood metabolites and proteins.

226

227 **MetBMI reflects gut microbiome profiles better than BMI**

228 Gut microbiome causally affects host obesity phenotypes in a mouse model³⁵ and obese human
229 individuals generally exhibit lower bacterial diversity and richness^{36,37}, while some meta-analyses of
230 human studies suggest an inconsistent relationship between gut microbiome and obesity^{38,39}. Given
231 our previous finding that the association between blood metabolites and bacterial diversity is
232 dependent on BMI¹⁹ and the current finding that the omics-based BMI models capture heterogeneous
233 metabolic health states of individuals, we hypothesized that the omics-based BMIs represent gut
234 microbiome diversity better than the measured BMI. As expected, while all BMI and omics-based
235 BMIs were significantly associated with different metrics of gut microbiome α -diversity (the number
236 of observed species, Shannon's index, and Chao1 index), the omics-based BMIs explained a larger

237 portion of the variance in the gut microbiome α -diversity than BMI (Shannon's index: ranging from
238 7.55% (ChemBMI) to 10.95% (MetBMI) compared to 6.62% (BMI); Fig. 4a, b). In particular,
239 MetBMI explained the largest portion of the variance, consistent with our previous observation that
240 plasma metabolomics showed a much stronger correspondence to gut microbiome structure than either
241 proteomics or clinical labs¹⁹.

242 We further examined the predictive power of gut microbiome profiles for the omics-based
243 BMI. We generated models classifying normal versus obese individuals using a random forest
244 classifier trained on gut microbiome 16S amplicon sequencing data. The gut microbiome-based
245 classifier for MetBMI categories showed significantly larger area under curve (AUC) in receiver
246 operator characteristic (ROC), sensitivity, and precision compared to these performance parameters of
247 the classifier for measured BMI categories ($P = 0.007$ (AUC), 0.007 (sensitivity), 0.019 (precision);
248 Fig. 4c, d). Therefore, these results suggest that the MetBMI model outperforms BMI even in its
249 capacity to reflect gut microbiome profiles.
250

251 **Metabolic health of the metabolically obese group was substantially improved following a** 252 **positive lifestyle intervention**

253 In the Arivale program, personalized lifestyle coaching was provided, resulting in significantly
254 positive clinical outcomes²². To investigate the corresponding longitudinal changes in omics-BMI
255 models, we utilized the available longitudinal measurements from a subcohort of 608 participants (see
256 Methods). Given the variability in time between data collection points, we estimated the mean
257 transition of the measured and omics-based BMIs using a linear mixed model (LMM). Consistent with
258 the previous analysis²², the mean BMI estimate for the overall cohort decreased during the program
259 (Fig. 5a). The decrease of MetBMI was larger than that of BMI while the decrease of ProtBMI was
260 minimal and even smaller than that of BMI (Fig. 5a), suggesting that plasma metabolomics are highly
261 responsive to weight loss, while proteomics (measured from the same blood draw) are resistant to the
262 same lifestyle coaching. Subsequently, we generated LMMs stratified by baseline BMI class, and
263 confirmed that a significant decrease in the mean BMI estimate was observed in the overweight and
264 obese BMI classes, but not in the normal BMI class (Fig. 5b). Concordantly, the mean estimates of
265 ProtBMI and ChemBMI exhibited negative changes over time in the overweight and obese BMI
266 classes, but not in the normal BMI class (Fig. 5b). However, the mean estimate of MetBMI exhibited a
267 significant decrease even in the normal BMI class (Fig. 5b), suggesting that metabolomics data may
268 capture information about the metabolic health response to lifestyle intervention, beyond baseline
269 BMI class or changes in BMI and other omics.

270 Since we revealed that multiple metabolic health states exist within the standard BMI classes
271 (Fig. 3), we further investigated the difference between misclassification strata based on the baseline
272 MetBMI class. In the normal baseline BMI class, while the mean BMI estimate remained constant in
273 both the matched and misclassified groups, the mean MetBMI estimate exhibited larger reduction in
274 the misclassified group than the matched group (Fig. 5c), suggesting that MONW participants
275 improved their metabolic health to a greater extent than metabolically healthy, normal-weight
276 (MHNW) participants. Likewise, in the obese baseline BMI class, while the decrease in the mean BMI
277 estimate was not significantly different between the matched and misclassified groups, the decrease in
278 the mean MetBMI estimate was larger in the matched group than in the misclassified group (Fig. 5d),
279 suggesting that MUO participants improved their metabolic health to a greater extent than MHO
280 participants. Altogether, these results suggest that metabolic health was substantially improved during
281 the program, in accordance with the baseline metabolomic state rather than the baseline BMI class.
282

283 **Plasma analyte correlation network in the metabolically obese group reverted back to normal** 284 **state following lifestyle intervention**

285 Finally, we explored longitudinal changes in plasma analytes correlation network, focusing on the
286 metabolically obese group. Based on the importance of baseline metabolomic state (Fig. 5c, d), we

287 first assessed relationships between each plasma analyte–analyte correlation and baseline MetBMI,
288 using their interaction term of a generalized linear model (GLM) for each analyte–analyte pair in 608
289 participants (See Methods). In this type of model, the statistical test assesses whether the relationship
290 between any two analytes is dependent on a third variable (in this case, baseline MetBMI). Among
291 608,856 pairwise relationships of plasma analytes, 91 analyte–analyte correlation pairs derived from
292 75 metabolites, 26 proteins, and 13 clinical lab tests were significantly modified by baseline MetBMI
293 (FDR < 0.05). Subsequently, we assessed longitudinal changes of the significant 91 pairs, using their
294 interaction term (days in the program) of a generalized estimating equation (GEE) for each pair in 184
295 metabolically obese participants (See Methods). Among the significant 91 pairs from the GLM
296 models, 14 analyte–analyte correlation pairs were significantly modified by days in the program (Fig.
297 6a). The significant 14 pairs were mainly derived from metabolites (16 metabolites, 3 clinical lab
298 tests). For instance, homoarginine is a recently discovered biomarker of CVD⁴⁰ and was a robust
299 positive predictor in the MetBMI and CombiBMI models (Fig. 2a and Supplementary Fig. 4a), while
300 phenyllactate (PLA) is a gut microbiome-derived phenylalanine derivative having antimicrobial
301 activity and natural antioxidant activity^{41,42}; and the positive correlation between homoarginine and
302 PLA was observed in metabolically obese group at baseline (Fig. 6b). However, this correlation in the
303 metabolically obese group was attenuated during the program (Fig. 6c), implying that some types of
304 plasma metabolic regulation specific to the metabolically obese group was improved during the
305 program.
306

307 Discussion

308 Obesity is a significant risk factor for many chronic diseases^{3–6}. The heterogeneous nature of human
309 health conditions, with variable manifestation ranging from metabolic abnormalities to cardiovascular
310 symptoms, calls for a deeper molecular characterization in order to optimize wellness and reduce the
311 current global epidemic of chronic diseases. In this study, we demonstrated that obesity profoundly
312 perturbs human physiology, as reflected across all the omics modalities studied. The key findings of
313 this study are: (1) machine learning-based multiomic BMI was a more reliable measure of metabolic
314 health than traditional BMI, while maintaining a high level of interpretability and intuitiveness
315 attributed to the original metric (Fig. 1–3); (2) metabolomic reflection of obesity exhibited the
316 strongest correspondence to gut microbiome profiles in all omics studied (Fig. 4); (3) the plasma
317 metabolomics exhibited stronger (and/or earlier) response to lifestyle coaching than measured BMI,
318 while the plasma proteomics exhibited a weaker (and/or more delayed) response (Fig. 5a, b); (4)
319 MONW (i.e., normal-weight by BMI but overweight or obese by MetBMI) participants exhibited a
320 greater improvement in metabolic state (but not in weight loss itself) in response to lifestyle coaching
321 compared to MHNW (i.e., normal-weight by both BMI and MetBMI) participants, and vice versa for
322 MHO and MUO participants (Fig. 5c, d); (5) dozens of analyte–analyte associations were modified by
323 the metabolic state, stressing the functional context dependence of many blood-measured analytes
324 (Fig. 6).

325 Multiple observational studies have explored obesity biomarkers. With regard to obesity, the
326 involvements of insulin/insulin-like growth factor (IGF) axis and chronic low-grade inflammation
327 have been discussed in the context of obesity-related disease risks^{5,6}, backed by robust findings about
328 the association of obesity with IGFBP1/2 (–), adipokines such as LEP (+), adiponectin (–), FABP4
329 (+), and ADM (+) and proinflammatory cytokines such as interleukin 6 (IL6, +)^{32,43}. Consistent with
330 these well-known changes, we observed the positive significant association of BMI with LEP,
331 FABP4, IL1RN, IL6, ADM, and insulin and the negative association with IGFBP1/2 and adiponectin
332 (Fig. 2c, d). Importantly, all these known biomarkers were incorporated into the omics-based BMI
333 models, and most of them emerged as the robustly retained strong components of the models (Fig. 2a
334 and Supplementary Fig. 4b, c). At the same time, we observed that RAGE explains a relatively small
335 proportion of the variance in BMI (Fig. 2c), while being a strong negative predictive variable in all ten
336 models of both ProtBMI and CombiBMI (Fig. 2a and Supplementary Fig. 4b). The importance of
337 soluble RAGE (sRAGE) has been gradually highlighted in T2DM and CVD⁴⁴, with several reports on

338 the negative association of sRAGE with BMI⁴⁵. Therefore, the omics-based BMI may not only reflect
339 obesity, but also reflect early transition to clinical manifestation of obesity-related chronic diseases.

340 Likewise, many epidemiological studies have revealed metabolomic biomarkers for
341 obesity^{46,47}. In line with the previous knowledge, we confirmed the positive significant association of
342 BMI with mannose, uric acid (urate), and glutamate and the negative correlation with asparagine and
343 glycine (Fig. 2b), and all of them were robustly incorporated into all ten models of MetBMI and
344 CombiBMI (Fig. 2a and Supplementary Fig. 4a). Furthermore, we observed that many lipids emerged
345 as the strong components in the MetBMI and CombiBMI models; in particular,
346 glycerophosphocholines (GPCs) are negative variable components while sphingomyelins (SMs) are
347 positive ones (Fig. 2a and Supplementary Fig. 4a), even though both have a phosphocholine group in
348 common. Although lipids were traditionally regarded as positively associated factors with obesity,
349 recent metabolomics studies have uncovered variable trends of fatty acid species; e.g., plasma
350 lysophosphatidylcholines (LPCs) are decreased in obese mice (high-fat diet model)⁴⁸, which was in
351 accordance with our analysis (e.g., LPC(18:1), described as 1-oleoyl-GPC(18:1), in Fig. 2b and
352 Supplementary Fig. 4b). However, because there are many combinations of acyl residues in lipids and
353 many potential confounding factors with obesity, a meta-analysis to systematically understand the
354 species-level lipid biomarkers for obesity is still challenging^{46,47}. Our approach, applying machine
355 learning to metabolomics data, addresses this challenge by automatically and systematically providing
356 a molecular signature of obesity, while reflecting the versatile and complex metabolite species.
357 Altogether, the omics-based BMI can be regarded as a multidimensional metric of obesity, possessing
358 molecular mechanistic information. Although targeted measurement of a specific metrics (e.g., body
359 composition) or biomarker provides a detailed, albeit narrow, view into the anthropometric parameters
360 of obesity, the blood multiomic measurements can therefore provide a broader view into its
361 heterogeneous molecular dimensions.

362 Recently, Cirulli and colleagues have reported a machine learning-based model of BMI
363 computed from blood metabolomics captured obesity-related phenotypes²⁰. Their main model
364 explained 39.1% of the variance in BMI, while our MetBMI model exhibited higher explained
365 variance (68.9%; Fig. 2b). Other than the difference in studied cohorts, the performance gap may also
366 result from the difference in modeling strategies. Cirulli and colleagues stringently selected 49
367 metabolites out of their metabolomics panel (1,007 metabolites) based on a pre-screening for
368 significant association with BMI, and subsequently applied a tenfold CV implementation of the Ridge
369 or LASSO method. In contrast, we used the LASSO method for feature selection, applying it to the
370 full metabolomics panel (766 metabolites). In addition to the increased number of modeled
371 metabolites, our higher performance may stem from the existence of metabolites which were critical
372 for BMI prediction but not significantly associated with BMI. In fact, our MetBMI model contained
373 many metabolites which were weakly associated with BMI but robustly retained across all ten models
374 (Fig. 2b and Supplementary Fig. 4a). At the same time, many of their 49 metabolites (14–20
375 metabolites among the 31–41 corresponding metabolites in our metabolomics panel) were also
376 retained in at least one of ten MetBMI models. Therefore, our strategy of feature selection through
377 machine learning may be preferable for predicting BMI from metabolomics.

378 A recent study that investigated multiomic changes in response to weight perturbation
379 demonstrated that some weight gain-associated blood signatures reverted during a subsequent weight
380 loss, while others persisted⁴⁹. Interestingly, we revealed that MetBMI is more responsive to lifestyle
381 intervention than BMI and ChemBMI, while ProtBMI is more resistant (Fig. 5a, b). Our analyses of
382 components in the omics-based BMI models (Fig. 2 and Supplementary Fig. 3b–e, 4) suggested that
383 various metabolites share a wider spread of the feature importance, while a small subset of proteins
384 (~5 proteins) predominantly reflects obesity profiles. Therefore, the effect of lifestyle coaching may
385 consist of small additive contributions in the short range rather than affecting the root cause. However,
386 longitudinal analysis with longer periods must be investigated to infer the physiological meaning of
387 the metabolomic responsiveness and the proteomic resistance. For instance, it is possible that ProtBMI
388 shows a delayed response to the same weight loss (over a span greater than a year measured presently;
389 Fig. 5a), indicating blood metabolites and proteins may be early and late responders to a lifestyle

390 intervention, respectively. If the difference between BMI and the omics-based BMIs remains constant
391 even after one year, it is reinforced that blood metabolites and proteins are more and less sensitive
392 factors of weight loss, respectively. In either scenario, the advantage of monitoring blood multiomics
393 during weight loss programs would be supported as a tool to keep participants motivated, since
394 lifestyle changes generally take longer to manifest as weight loss. At the same time, long-term
395 maintenance of the improvement is an important challenge for lifestyle interventions. Although there
396 is variability between studies, a previous study indicated that only $\approx 20\%$ of overweight individuals
397 successfully maintained weight loss⁵⁰, raising the possibility that BMI and the other omics-based
398 BMIs revert back to the baseline after one year of the Arivale program. Nevertheless, previous studies
399 revealed that lifestyle interventions had benefits in preventing diabetes incidence as long as 20 years
400 post-intervention even if body weight was regained^{51,52}, implying that the observed larger
401 improvement of MetBMI compared to BMI may persist in the longer term. Hence, deeper
402 investigations on our findings are required, especially the long-term dynamics of the MetBMI and
403 ProtBMI responses, which may provide a foothold to develop a scientific strategy to maintain
404 metabolic health in the long term. In addition to the studied time period, there are additional
405 limitations to be noted in this study. This study was not designed as a randomized control trial, and we
406 cannot strictly evaluate the effectiveness of the lifestyle intervention (e.g., improvement in obese
407 group due to the regression-toward-the-mean effect). As an observational study derived from a
408 consumer facing cohort, generalizability of the findings may be limited. Our measurements did not
409 cover all biomolecules in blood; in particular, proteomics was based on targeted Olink panels. Hence,
410 our findings on metabolomic and proteomic states are restricted to the studied space. Nevertheless,
411 this study will serve as a valuable resource for robustly characterizing metabolic health from the blood
412 and identifying actionable targets for health interventions.
413

414 **Methods**

415 **Study cohort**

416 This study relied on a cohort consisting of over 5,000 individuals who participated in a wellness
417 program offered by a currently closed commercial company (Arivale Inc., Washington, USA). We
418 collected personal, dense, dynamic data (PD3) clouds on individuals in this program between 2015–
419 2019. An individual was eligible for enrollment if the individual was over 18 years old, not pregnant,
420 and a resident of any US state except New York; participants were primarily recruited from
421 Washington, California, and Oregon. During the Arivale program, each participant was provided
422 personalized lifestyle coaching via telephone by registered dietitians, certified nutritionists, or
423 registered nurses (see the previous paper²² for details). Participants also had access to their clinical
424 data via an online data dashboard. This study was conducted with deidentified data of the participants
425 who had consented to the use of their anonymized data in research. All procedures were approved by
426 the Western Institutional Review Board (WIRB) with Institutional Review Board (IRB) (Study
427 Number: 20170658 at Institute for Systems Biology and 1178906 at Arivale).

428 In this study, to confidently compare the association between Body Mass Index (BMI) and
429 host phenotypes across multiomics, we limited our study cohort to only the participants whose
430 datasets contained (1) all main omic measurements (metabolomics, proteomics, clinical lab tests) from
431 the same first blood draw, (2) genetic information, and (3) a baseline BMI measurement within ± 1.5
432 month from the first blood draw. We also eliminated “outlier” participants whose baseline BMI was
433 beyond ± 3 standard deviations from mean in the cohort distribution. The final cohort consists of 1,277
434 participants, whose demographics (Supplementary Fig. 1) were confirmed to be consistent with the
435 overall Arivale demographics previously reported in our studies^{19,22–26}.

436

437 **Data collection and PD3 clouds**

438 The PD3 clouds consist of human genomes, longitudinal measurements of metabolomics, proteomics,
439 clinical lab tests, gut microbiomes, and wearable devices, and health/lifestyle questionnaires.
440 Peripheral venous blood draws for all measurements were performed by trained phlebotomists at
441 LabCorp (Laboratory Corporation of America Holdings, North Carolina, USA) or Quest (Quest
442 Diagnostics, New Jersey, USA) service centers. For some participants, saliva was also sampled at
443 home to measure analytes such as diurnal cortisol and dehydroepiandrosterone (DHEA) using a
444 standardized kit (ZRT Laboratory, Oregon, USA). Likewise, stool samples for gut microbiome
445 measurements were obtained by participants at home using a standardized kit (DNA Genotek, Inc.,
446 Ottawa, Canada).

447 **– Genome**

448 DNA extracted from whole blood underwent whole genome sequencing or single-nucleotide
449 polymorphisms (SNP) microarray genotyping. Genetic ancestry was calculated with principal
450 components (PCs) using a set of $\sim 100,000$ ancestry-informative SNP markers as described
451 previously²². Polygenic risk scores were constructed using publicly available summary
452 statistics from published genome-wide association studies (GWAS) as described previously²⁴.

453

454 **– Blood-measured omics**

455 Metabolomic data was generated by Metabolon, Inc. (North Carolina, USA), using ultra-high-
456 performance liquid chromatography-tandem mass spectrometry (UHPLC-MS/MS) for plasma
457 derived from whole blood samples. Proteomic data was generated using proximity extension
458 assay (PEA) for plasma derived from whole blood samples with several Olink Target panels
459 (Olink Proteomics, Uppsala, Sweden), and measurements with the Cardiovascular II,
460 Cardiovascular III and Inflammation panels were used in this study since the other panels
461 were not necessarily applied to all samples. All clinical laboratory tests were performed by
462 LabCorp or Quest in a Clinical Laboratory Improvement Amendments (CLIA)-certified lab,

463 and measurements by LabCorp were selected in this study to eliminate potential differences
464 between vendors. In this study, analytes missing in more than 10% of the baseline samples
465 were removed from each omic dataset, and observations missing in more than 10% of the
466 remaining analytes were further removed. The final filtered metabolomics, proteomics, and
467 clinical labs consist of 766 metabolites, 274 proteins, 71 clinical lab tests, respectively
468 (Supplementary Data 1).
469

470 – Gut microbiome

471 Gut microbiome data was generated based on 16S amplicon sequencing of the V4 region
472 using a MiSeq sequencer (Illumina, Inc., California, USA) for DNA extracted from stool
473 samples, as previously described²⁵. All samples were first rarefied to an even sampling depth
474 of 38,813, the minimum number of reads per sample in the cohort. Gut microbiome α -
475 diversity was calculated at the amplicon sequence variant (ASV) level using Shannon’s index
476 calculated by $H = -\sum_{i=1}^S p_i \ln p_i$, where p_i is the proportion of species i to the total number
477 of species S in a community represented by ASVs or using Chao1 diversity score calculated
478 by $S_{\text{Chao1}} = S_{\text{obs}} + \frac{n_1^2}{2n_2}$, where S_{obs} is the number of observed ASVs, n_1 is the number of
479 singletons (ASVs captured once), and n_2 is the number of doubletons (ASVs captured twice).
480

481 – Anthropometrics, saliva-measured analytes, daily physical activity measures

482 Anthropometrics including weight, height, and waist circumference and blood pressure were
483 measured at the time of blood draw and also reported by participants, which generated diverse
484 timing and number of observations depending on each participant. The measured weight and
485 height simultaneously generated BMI measurements with dividing the weight by the squared
486 height. Measurements of saliva samples were performed in the testing laboratory of ZRT
487 Laboratory. Daily physical activity measures such as heart rate, moving distance, number of
488 steps, burned calories, floors climbed, and sleep quality were tracked using the Fitbit
489 wearable device (Fitbit, Inc., California, USA). To manage variations between days, monthly
490 averaged data was used for these daily measures. In this study, the baseline measurement for
491 these longitudinal measures was defined with the closest observation to the first blood draw
492 per participant and data type, and each dataset was eliminated from analyses when its baseline
493 measurement was beyond ± 1.5 month from the first blood draw.
494

495 Generation of omics-based BMI models

496 For each omic dataset, missing values were first imputed with a random forest (RF) algorithm using
497 Python missingpy (version 0.2.0) library corresponding to MissForrest in R⁵³. For sex-stratified
498 models, the datasets after imputation were divided into sex-dependent datasets. Each value was
499 subsequently standardized with the Z-score using mean and standard deviation per analyte. Then, ten
500 iterations of least absolute shrinkage and selection operator (LASSO) modeling with tenfold cross-
501 validation (CV) were performed for the log-transformed BMI and the processed omic datasets using
502 Python scikit-learn (version 0.22.1) library. Training and testing datasets were generated by splitting
503 participants into ten sets with one set as a “testing dataset” and the remaining nine sets as a “training
504 dataset”, and iterating all combinations over those ten sets; i.e., overfitting was controlled using an
505 internal tenfold CV in each “training dataset”. Consequently, this procedure generated a “testing
506 dataset”-derived BMI prediction value for each participant and ten fitted models for each omics.
507 Model performance was conservatively estimated by the R^2 score from out-of-sample predictions.
508 Pearson’s r was calculated using measured and predicted BMI values for the entire cohort.

509 For the LASSO-modeling iterations while dropping the strongest variable, the generation of
510 ten LASSO models was repeated as the same as the above except for eliminating the strongest

511 variable analyte from the dataset at the end of each iteration. The strongest variable was defined as the
512 variable that was retained across ten models and had the highest absolute value of mean β -coefficient.

513 For longitudinal models, the standardization distribution and training dataset were restricted
514 to the baseline measurements from all 1,277 participants and only one LASSO model with tenfold CV
515 was generated per sex-stratified cohort, because those measurements were minimally affected by
516 lifestyle coaching and each participant had a different number of observations.
517

518 **Generation of obesity-classifying models**

519 To compare the ability of the gut microbiome to accurately distinguish between obese and normal
520 weight individuals across different omics measurements of BMI (Fig. 4c, d), ASVs were collapsed
521 into species, genus, family, order, class, and phylum, respectively, and merged into a single dataframe
522 using the SILVA database (version 132). This dataframe served as input for a RF classifier
523 implemented with Python scikit-learn library. Briefly, two classifiers were trained on taxon
524 abundances, one predicting whether an individual is obese based on observed BMI, and one predicting
525 whether an individual is obese based on metabolomics-based BMI. Both models were constructed
526 using RF with a fivefold CV scheme. In this approach, 80% of the data is used for training while the
527 remaining 20% is used as a testing set. This process is repeated fivefold where each participant serves
528 as part of the testing set once. Performance for each of the classifiers was then assessed by averaging
529 the performance across the five testing sets.
530

531 **Longitudinal analysis for BMI and omics-based BMIs**

532 For longitudinal analyses, 608 participants, whose datasets contained more than two time-series
533 datasets of both BMI and omics during 18 months after enrollment, were further extracted from the
534 study cohort of 1,277 participants. To estimate the mean transition of the measured and omics-based
535 BMIs, a linear mixed model (LMM) for the rate of change in each measured BMI or sex-stratified
536 LASSO models-predicted BMI was generated with random intercepts for participants and random
537 slopes for days in the program, following the previous approach²². As the fixed effects regarding time,
538 linear regression splines with knots at 0, 6, 12, and 18 months were applied to days in program to fit
539 time as a continuous variable rather than a categorical variable because the timing of data collection
540 was different between the participants, allowing for differences in the trajectory of changes throughout
541 the program. In addition to the linear regression splines for days in the program, each LMM included
542 sex, baseline age, ancestry PCs, and meteorological seasons as fixed effects to adjust potential
543 confounding effects. For the baseline BMI class-stratified LMMs, the interaction terms between the
544 categorical baseline BMI-based class and the linear regression splines for time were further added. All
545 LMMs were modeled using Python statsmodels (version 0.11.1) library. Of note, the underweight
546 participants were eliminated in the LMMs stratified with baseline BMI class because the sample size
547 was too small for convergence.
548

549 **Plasma analyte correlation network analysis**

550 In advance, outlier values which were beyond ± 3 standard deviations from mean in the longitudinal
551 cohort distribution of 608 participants were eliminated from the dataset per analyte, and seven clinical
552 lab tests which became almost invariant across the participants were eliminated from analyses,
553 allowing convergence in the following modeling. Against each analyte, values were converted with a
554 transformation method producing the lowest skewness (e.g., no transformation, the logarithm
555 transformation for right skewed distribution, the square root transformation with mirroring for left
556 skewed distribution) and standardized with the Z-score using mean and standard deviation.

557 Against 608,856 pairwise combinations of the analytes (766 metabolites, 274 proteomics, 64
558 clinical lab tests), generalized linear models (GLMs) for the baseline measurements of 608
559 participants were independently generated with the Gaussian distribution and identity link function

560 using Python statsmodels library. Each GLM constitutes of an analyte as dependent variable, another
561 analyte and the baseline MetBMI as independent variables with their interaction term, and sex,
562 baseline age, and ancestry PCs as covariates. The significant analyte–analyte correlation pairs
563 modified by the baseline MetBMI were obtained based on the β -coefficient (two-sided t -test) of the
564 interaction term between independent variables in GLM, while correcting the multiple-hypothesis
565 testing with the Benjamini–Hochberg method (false discovery rate (FDR) < 0.05).

566 Against the significant 91 pairs (75 metabolites, 26 proteomics, 13 clinical lab tests) from the
567 GLM analysis, generalized estimating equations (GEEs) for the longitudinal measurements of 184
568 metabolically obese participants were independently generated with the exchangeable covariance
569 structure using Python statsmodels library. Each GEE constitutes of an analyte as dependent variable,
570 another analyte and days in the program as independent variables with their interaction term, and sex,
571 baseline age, ancestry PCs, and meteorological seasons as covariates. The significant analyte–analyte
572 correlation pairs modified by days in the program were obtained based on the β -coefficient (two-sided
573 t -test) of the interaction term between independent variables in GEE, while correcting the multiple-
574 hypothesis testing with the Benjamini–Hochberg method (FDR < 0.05).
575

576 **Statistical analysis**

577 All data preprocessing and ordinary least squares (OLS) regression analyses were performed using
578 Python NumPy (version 1.18.1), pandas (version 1.0.3), SciPy (version 1.4.1) and statsmodels
579 libraries. Only the baseline datasets were utilized in regression analyses, and each numeric variable
580 was scaled and centered in advance. When assessing available numeric physiological features and
581 obesity-related health markers in the PD3 clouds, the baseline dataset of each metric variable was also
582 preprocessed with the elimination steps for outliers and invariant variables and the conversion step for
583 skewness reduction, as same as those described in the above subsection except for the basis of whole
584 study cohort distribution. Relationships of the preprocessed numeric physiological features with the
585 measured or omics-based BMIs (Fig. 1c) were independently assessed using OLS linear regression
586 with the log-transformed measured or omics-based BMI as dependent variable and sex, age, and
587 ancestry PCs as covariates, while correcting the multiple-hypothesis testing with the Benjamini–
588 Hochberg method (FDR < 0.05). Relationships between BMI and analytes which were retained in at
589 least one of ten LASSO models (210 metabolites, 75 proteins, 42 clinical lab tests) (Fig. 2b–d) were
590 independently assessed using OLS linear regression with the log-transformed BMI as dependent
591 variable and sex, age, and ancestry PCs as covariates, while correcting the multiple-hypothesis testing
592 with the Benjamini–Hochberg method (FDR < 0.05). Differences in the BMI-associated features and
593 obesity-related health markers between the matched and misclassified groups in the normal or obese
594 BMI class (Fig. 3c and Supplementary Fig. 5a) were independently assessed using OLS linear
595 regression with sex, age, and ancestry PCs as covariates (sex and ancestry PCs as covariates for the
596 regression of age). Misclassification distribution in hierarchical clustering (Figs. 3d, e and
597 Supplementary Fig. 5b, c) was assessed using Fisher’s exact tests with the Bonferroni correction
598 (family-wise error rate (FWER) < 0.05). Relationships between measured or omics-based BMI and α -
599 diversity metrics (Fig. 4a, b) were independently assessed using OLS linear regression with α -
600 diversity as dependent variable and sex, age, and ancestry PCs as covariates. Difference in classifier
601 performance parameters (Fig. 4d) was assessed using Student’s t -test. All statistical tests were
602 performed using a two-sided hypothesis.
603

604 **Data visualization**

605 Almost all results were visualized using Python matplotlib (version 3.2.1) and seaborn (version
606 0.10.1) libraries. Data were summarized as the mean \pm standard error of the mean (s.e.m.), the mean
607 with 95% confidence interval (CI), or the boxplot (median: center line, 95% CI around median: notch,
608 [Q_1 , Q_3]: box limits, [$\max(\text{minimum value}, Q_1 - 1.5 \times \text{IQR})$, $\min(\text{maximum value}, Q_3 + 1.5 \times \text{IQR})$]:
609 whiskers, where Q_1 , Q_3 , and IQR are the 1st quartile, the 3rd quartile, and the interquartile range,
610 respectively), as indicated in each figure legend. For presentation purpose, s.e.m. and CI were

611 simultaneously calculated during visualization using seaborn barplot or boxplot (utilizing matplotlib)
612 application programming interface (API) with default setting (1,000 times bootstrapping or a
613 Gaussian-based asymptotic approximation, respectively). The OLS linear regression line with 95% CI
614 was simultaneously generated during visualization using seaborn lmlot API with default setting
615 (1,000 times bootstrapping). Hierarchical clustering was simultaneously performed during
616 visualization using seaborn clustermap API (utilizing SciPy library) with the Ward's linkage method
617 for Euclidean distance. The plasma analyte correlation network was visualized with a circo plot using
618 R circlize (version 0.4.11) package⁵⁴.
619

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- 729

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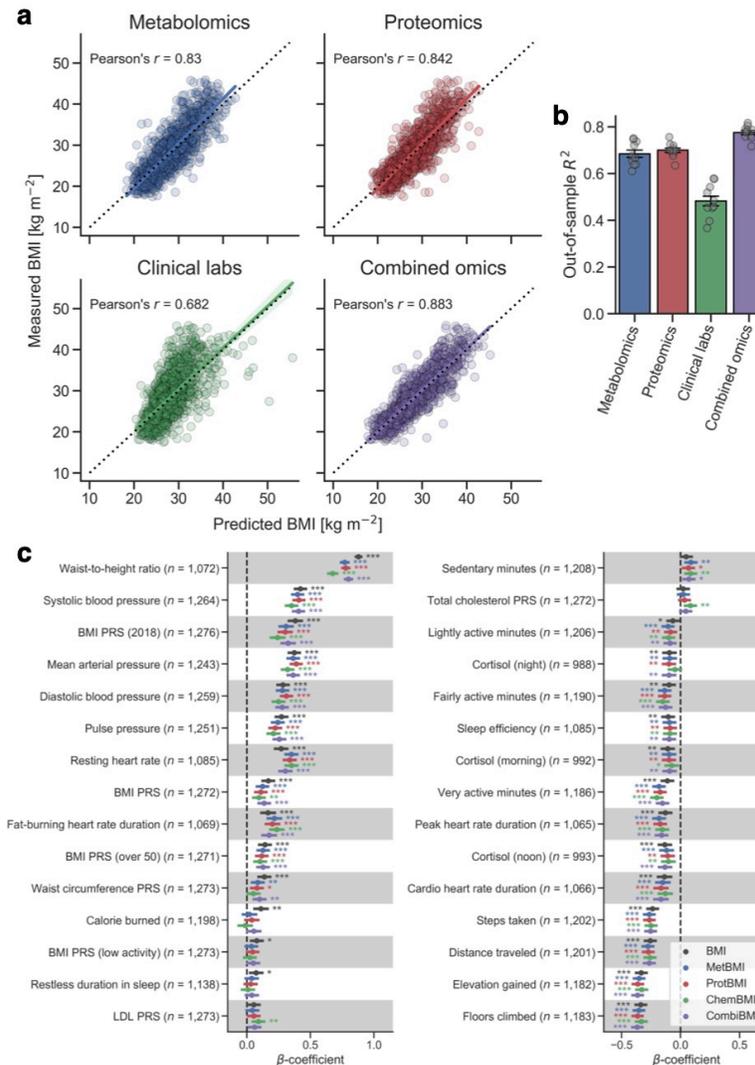
739 **Author Contribution**

740 K.W., T.W., L.H., N.D.P., and N.R. conceptualized the study. K.W., T.W., A.Z., N.D.P., and N.R.
741 participated in the study design. K.W., T.W., C.D., B.L., and N.R. performed data analysis and figure
742 generation. J.J.H., J.C.L., C.D., A.T.M., and L.H. assisted in results interpretation. J.C.L. and A.T.M.
743 managed the logistics of data collection and integration. K.W., T.W., and N.R. were the primary
744 authors of the paper, with contributions from all other authors. All authors read and approved the final
745 manuscript.
746

747 **Competing Interests**

748 J.J.H. has received grants from Pfizer and Novartis for research unrelated to this study. All other
749 authors declare no competing interests.
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751 **Figures**



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Figure 1. Plasma multiomics captures 48–78% of the variance in BMI.

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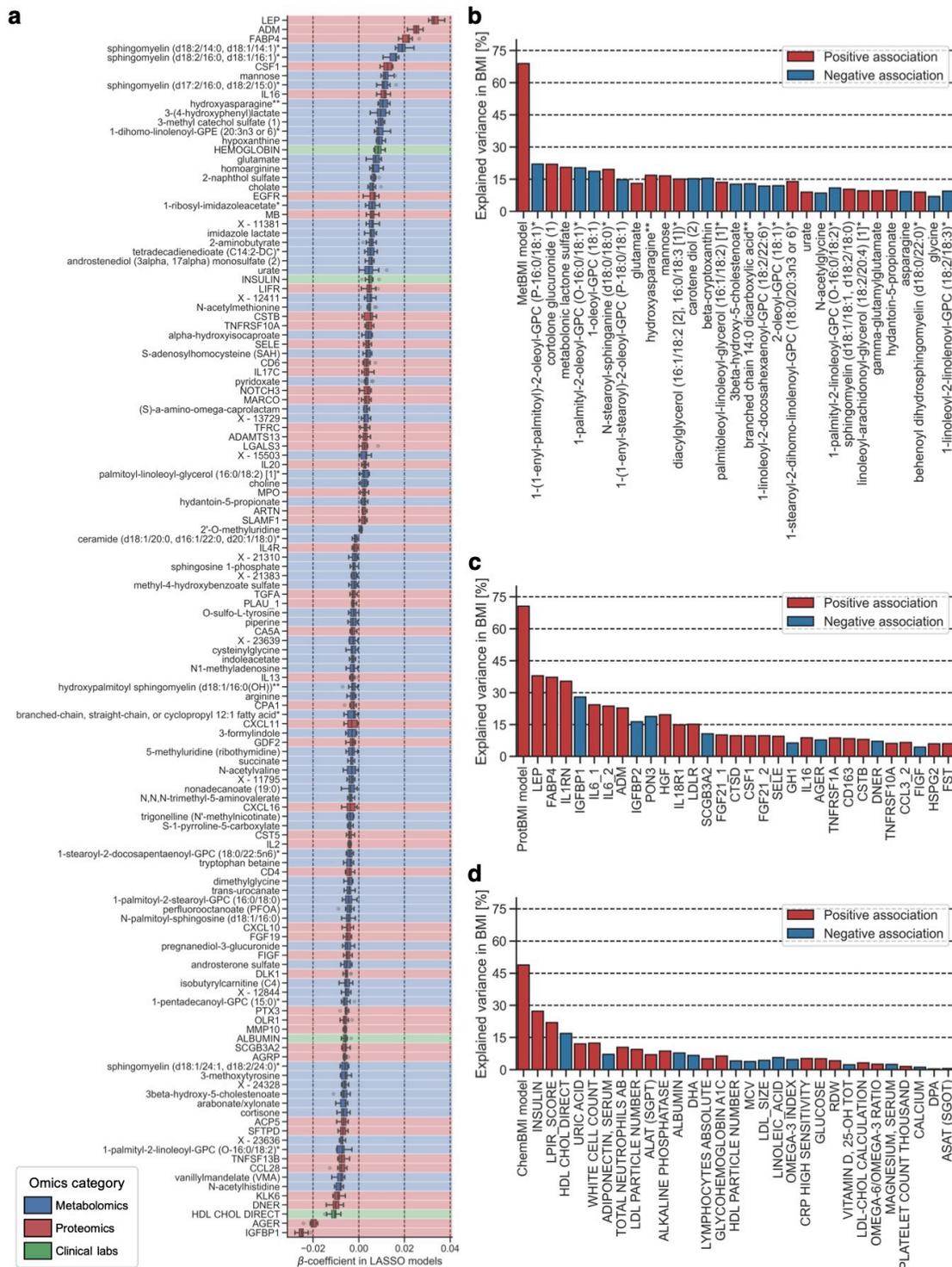
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a Scatterplot of measured Body Mass Index (BMI) versus predicted BMI using least absolute shrinkage and selection operator (LASSO) with tenfold cross-validation (CV). The solid line in each panel is the ordinary least squares (OLS) linear regression line with 95% confidence interval (CI), and the dotted line is measured BMI = predicted BMI. $n = 1,277$ participants. **b** Mean out-of-sample R^2 across the tenfold CV for each omics. Data: mean \pm s.e.m., $n = 10$ LASSO models. **c** β -coefficients for numeric physiological feature in each OLS linear regression model with BMI or omics-based BMI as dependent variable and sex, age, and ancestry principal components (PCs) as covariates. All presented 30 features are significantly associated with at least one of BMI or omics-based BMIs in the Benjamini–Hochberg method (false discovery rate (FDR) < 0.05 ; $*P < 0.05$, $**P < 0.01$, $***P < 0.001$). BMI: measured BMI, MetBMI: metabolomics-based BMI, ProtBMI: proteomics-based BMI, ChemBMI: clinical chemistries-based BMI, CombiBMI: combined omics-based BMI, PRS: polygenic risk score, LDL: low-density lipoprotein, n : the number of participants. Data: β -coefficient with 95% CI.



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Figure 2. Omics-based BMI captures the variance in BMI better than any single analyte.

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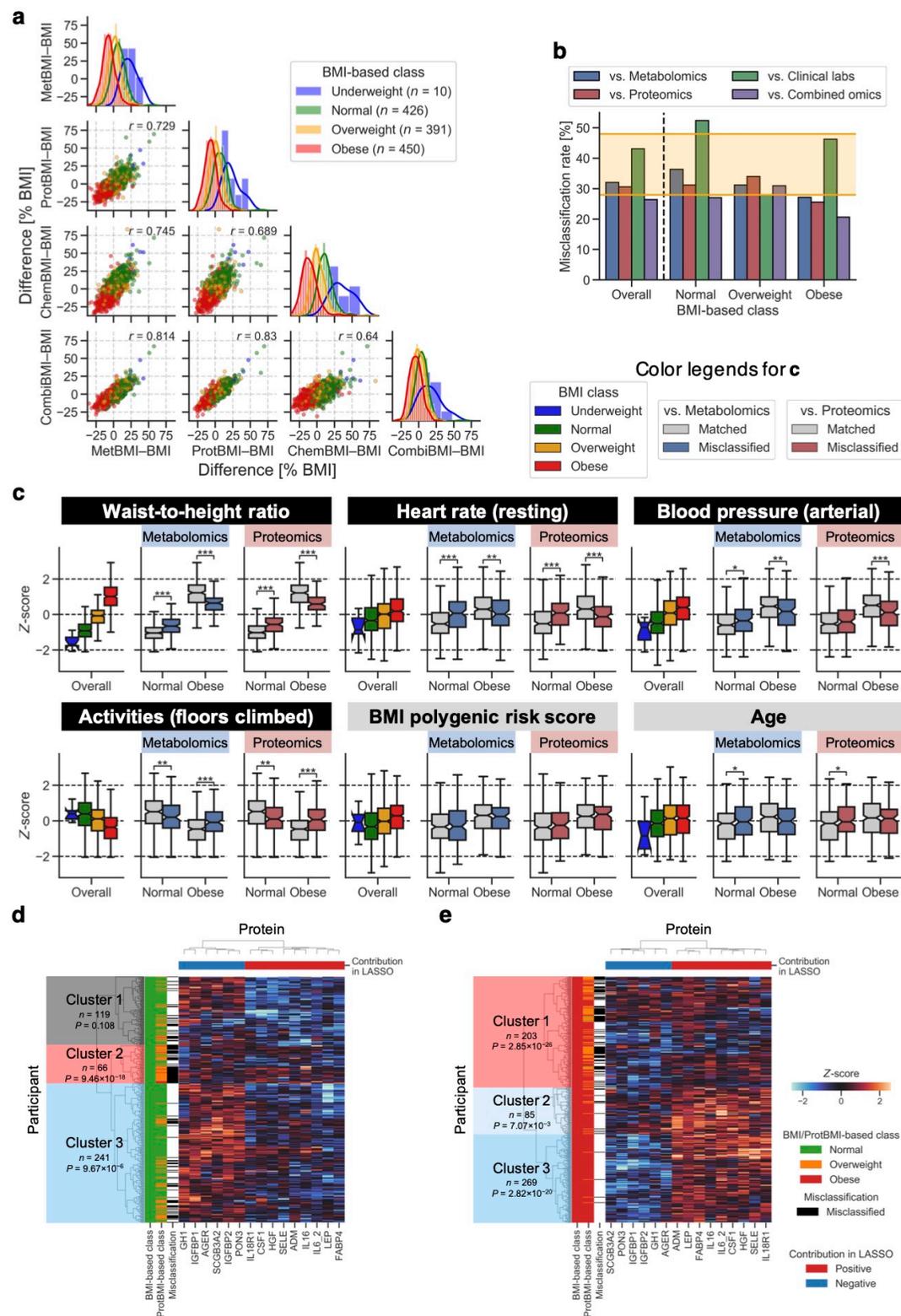
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a β -coefficient estimates for the variables that were retained across all ten combined omics-based Body Mass Index (BMI) models (132 analytes). Color of each row corresponds to the analyte category (blue: 77 metabolites, red: 51 proteins, green: 4 clinical lab tests). Data: median (center line), [Q_1 , Q_3] (box limits), [$\max(\text{minimum value}, Q_1 - 1.5 \times \text{IQR})$, $\min(\text{maximum value}, Q_3 + 1.5 \times \text{IQR})$] (whiskers), where Q_1 , Q_3 , and IQR are the 1st quartile, the 3rd quartile, and the interquartile range,

775 respectively; $n = 10$ least absolute shrinkage and selection operator (LASSO) models. **b–d** Percentage
776 of variance in BMI explained by each metabolite (**b**), protein (**c**), or clinical lab test (**d**). BMI was
777 independently regressed to each analyte which was retained in at least one of ten LASSO models (210
778 metabolites, 75 proteins, 42 clinical lab tests). The strongest 30 analytes among the analytes
779 significantly associated with BMI (180 metabolites, 63 proteins, 30 clinical lab tests) are presented.
780 Significance was assessed using ordinary least squares (OLS) linear regression with sex, age, and
781 ancestry principal components (PCs) as covariates, while correcting for multiple-hypothesis testing
782 with the Benjamini–Hochberg method (false discovery rate (FDR) < 0.05). Each omics-based BMI is
783 included for comparison (MetBMI: metabolomics-based BMI, ProtBMI: proteomics-based BMI,
784 ChemBMI: clinical chemistries-based BMI).
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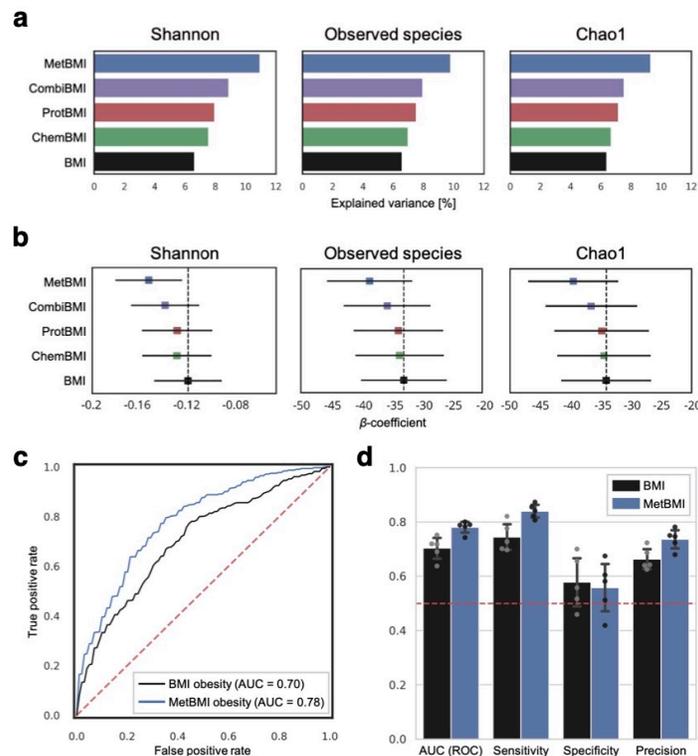
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Figure 3. Metabolic heterogeneity within standard BMI classes underlies the high rate of misclassification.

789 **a** Scatterplot and distribution of difference between Body Mass Index (BMI) and omics-based BMI.
790 BMI: measured BMI, MetBMI: metabolomics-based BMI, ProtBMI: proteomics-based BMI,
791 ChemBMI: clinical chemistries-based BMI, CombiBMI: combined omics-based BMI, r : Pearson's
792 correlation coefficient, n : the number of participants. The line in histogram panels indicates the kernel
793 density estimate. **b** Misclassification rate of BMI-based class. Range of previously reported
794 misclassification rate^{30,31} is highlighted with orange-colored lines. Note that the underweight BMI
795 class is not presented due to small sample size, and its misclassification rate is 80% against combined
796 omics and 100% against the others. **c** Comparison of BMI-associated feature between the matched and
797 misclassified groups in the normal or obese BMI class. Data: median (center line), 95% confidence
798 interval (CI) around median (notch), [Q_1 , Q_3] (box limits), [$\max(\text{minimum value}, Q_1 - 1.5 \times \text{IQR})$,
799 $\min(\text{maximum value}, Q_3 + 1.5 \times \text{IQR})$] (whiskers), where Q_1 , Q_3 , and IQR are the 1st quartile, the 3rd
800 quartile, and the interquartile range, respectively. $*P < 0.05$, $**P < 0.01$, $***P < 0.001$ according to
801 ordinary least squares (OLS) linear regression with sex, age, and ancestry principal components (PCs)
802 as covariates (sex and ancestry PCs as covariates for the regression of age). **d**, **e** Heatmap with
803 hierarchical clustering of the normal (**d**) and obese (**e**) BMI class using proteomics data. The strongest
804 15 proteins among the analytes retained across more than eight ProtBMI models and significantly
805 associated with BMI were used as variables. Z -score was calculated from the overall population. n : the
806 number of participants, P : the adjusted P -value in multiple-hypothesis testing for the misclassification
807 distribution in each cluster using two-sided Fisher's exact tests with the Bonferroni correction.
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Figure 4. Metabolomics-based BMI reflects gut microbiome profiles better than BMI.

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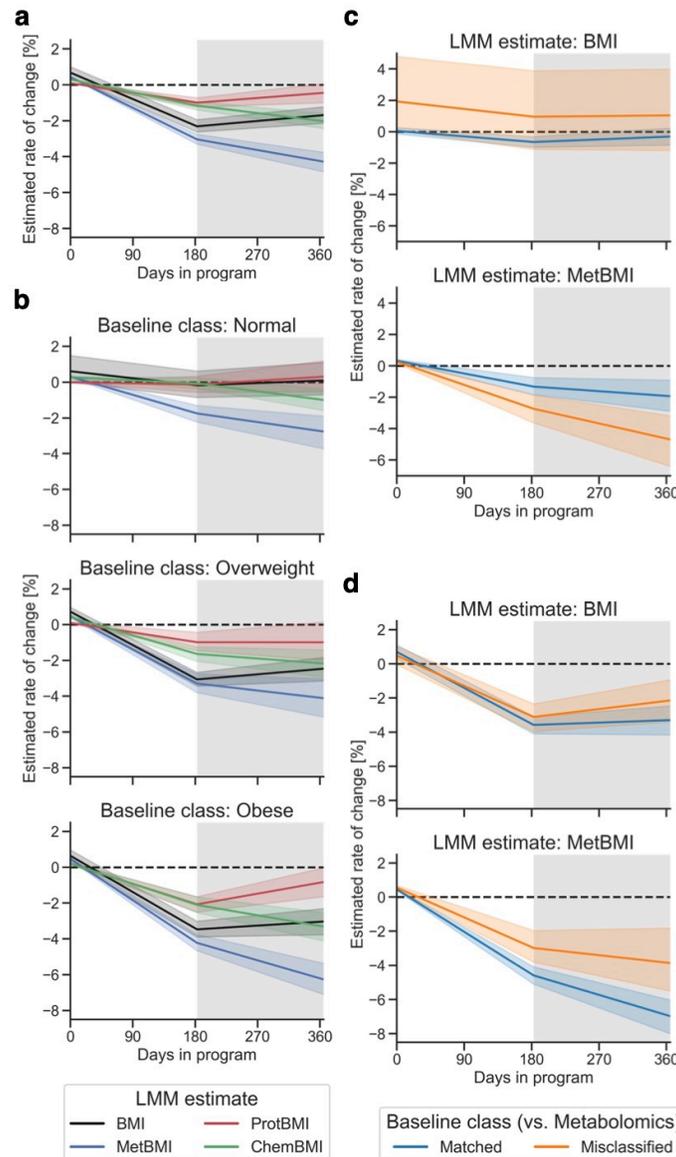
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a Percentage of variance in gut microbiome α -diversity that is explained by each Body Mass Index (BMI) or omics-based BMI. BMI: measured BMI, MetBMI: metabolomics-based BMI, ProtBMI: proteomics-based BMI, ChemBMI: clinical chemistries-based BMI, CombiBMI: combined omics-based BMI. **b** β -coefficient for BMI or omics-based BMI in each ordinary least squares (OLS) linear regression model with α -diversity as dependent variable and sex, age, and ancestry principal components (PCs) as covariates. The dashed line indicates the β -coefficient estimate for BMI. Data: β -coefficient with 95% confidence interval (CI). **c** Receiver operator characteristic (ROC) curve of a gut microbiome-based model classifying participants to the normal vs. obese class. Gut microbiome 16S ribosomal RNA datasets were used for generating the random forest (RF) classifier with fivefold cross-validation (CV). Each ROC curve indicates the average curve across five RF models. The red dashed line indicates a random classification line. AUC: area under curve. **d** Comparison of AUC of ROC curve, sensitivity, specificity, and precision between the classifying models of BMI and MetBMI. Each performance parameter was calculated as the mean out-of-sample value across the fivefold CV. Data: mean with 95% CI, $n = 5$ RF models.



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Figure 5. Metabolic health of the metabolically obese group was substantially improved following a positive lifestyle intervention.

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a Longitudinal change in Body Mass Index (BMI) or omics-based BMI for overall cohort. Rate of change in BMI and omics-based BMIs was estimated using each linear mixed model (LMM) with random intercepts for participants and random slopes for days in the program (see Methods). $n = 608$ participants. **b** Longitudinal change in BMI or omics-based BMI for each baseline BMI-based class. Rate of change in BMI and omics-based BMIs was estimated using each baseline BMI-based class-stratified LMM with random intercepts for participants and random slopes for days in the program. $n = 222$ (Normal), 185 (Overweight), 196 (Obese) participants. **c, d** Longitudinal change in BMI or metabolomics-based BMI of the participants misclassified with the normal (**c**) or obese (**d**) BMI class. $n = 156$ (Normal, Matched), 66 (Normal, Misclassified), 151 (Obese, Matched), 45 (Obese, Misclassified) participants. **a–d** The dashed line and gray shading correspond to the baseline value of each estimate and the 2nd period of linear regression spline for time, respectively. BMI: measured BMI, MetBMI: metabolomics-based BMI, ProtBMI: proteomics-based BMI, ChemBMI: clinical chemistries-based BMI. Data: mean with 95% confidence interval (CI).

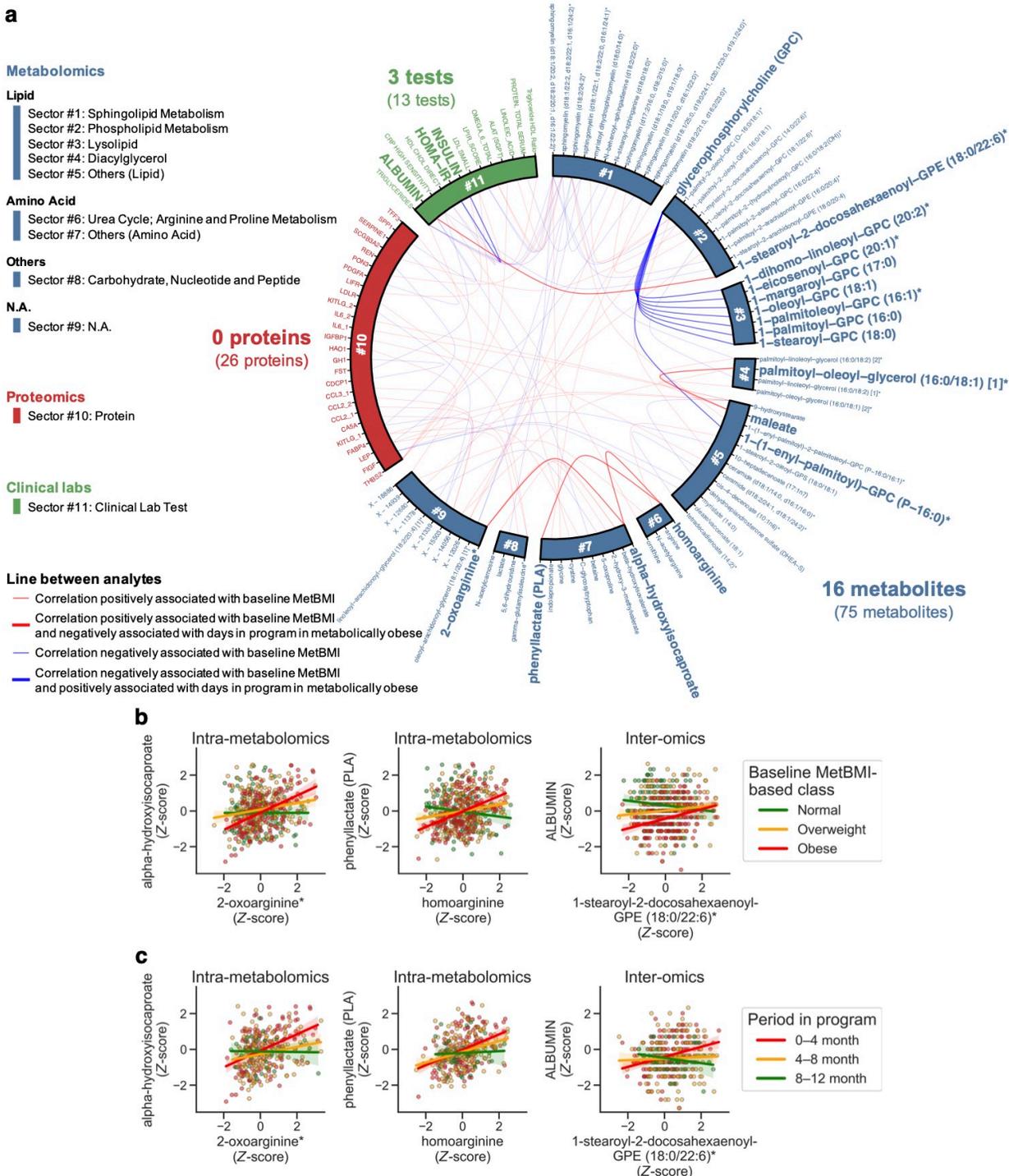


Figure 6. Plasma analyte correlation network in the metabolically obese group reverted back to normal state following lifestyle intervention.

a Circos plot of cross-omic interactions modified by metabolomics-based Body Mass Index (MetBMI) and days in the program. Among 608,856 pairwise relationships of plasma analytes (766 metabolites, 274 proteomics, 64 clinical lab tests) from 608 participants, 91 analyte-analyte pairs significantly modified by the baseline MetBMI are presented (75 metabolites, 26 proteomics, 13 clinical lab tests), whose significance was assessed using their interaction term in each generalized linear model (GLM);

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851 see Methods) while correcting the multiple-hypothesis testing with the Benjamini–Hochberg method
852 (false discovery rate (FDR) < 0.05). Among the significant 91 pairs from 184 metabolically obese
853 participants, 14 analyte–analyte pairs significantly modified by days in the program are highlighted by
854 line width and label font size (16 metabolites, 3 clinical lab tests), whose significance was assessed
855 using their interaction term in each generalized estimating equation (GEE; see Methods) while
856 correcting the multiple-hypothesis testing with the Benjamini–Hochberg method (FDR < 0.05). **b, c**
857 Representative examples of the analyte–analyte pair significantly modified by the baseline MetBMI
858 (**b**) or days in the program (**c**) in **a**. The solid line in each panel is the ordinary least squares (OLS)
859 linear regression line with 95% confidence interval (CI). $n = 530$ (**b**, Intra-metabolomics (left)), 553
860 (**b**, Intra-metabolomics (right)), 566 (**b**, Inter-omics) participants; $n = 329$ (**c**, Intra-metabolomics
861 (left)), 344 (**c**, Intra-metabolomics (right)), 353 (**c**, Inter-omics) measurements from 184 metabolically
862 obese participants. Of note, data points outside of plot range are trimmed in these presentations.
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